HawkesProcesses.jl

Dean Markwick

May 22, 2020

Abstract

HawkesProcesses.jl is a Julia package for using Hawkes processes to model event time data. This package provides functions to plot the intensity, calculate that likelihood and fit the parameters using a latent variable Bayesian algorithm from [Ros16] and extending it to the generic case.

1 Introduction

A Hawkes process is a point process where the intensity function is self-exciting. The intensity at time t given the history of the process, H_t can be written as

$$\lambda(t \mid H_t) = \mu(t) + \kappa \sum_{t_i < t} g(t - t_i), \tag{1}$$

where μ is the background rate, κ is a constant parameter and g(t) is some positive function.

2 Likelihood

For a collection of event times t_i, \ldots, t_n in the time interval [0, T] the log likelihood can be written as

$$\mathcal{L} = \sum_{i=1}^{n} \log \lambda(t_i \mid H_{t_i}) - \int_0^T \lambda(t \mid H_t) \mathrm{d}t,$$

which is also the general log likelihood for a point process from [DVJ03]. By inserting the Hawkes process intensity function into this equation we can simplify it down.

The integral of the intensity can be expanded as

$$\int_{0}^{T} \lambda(t \mid H_{t}) dt = \int_{0}^{T} \mu(t) dt + \kappa \int_{0}^{T} \sum_{t_{j} < t} g(t - t_{j})$$
$$= M(T) + \kappa \sum_{i=1}^{n} \left(G(T - t_{i}) - G(0) \right),$$

where $M(T) = \int_0^T \mu(t) dt$ and $G(t) = \int g(t) dt$. As we are forcing g to be a probability density defined on 0 to T, G is just the distribution function and thus G(0) = 0. Thus the final likelihood can be written as

$$\mathcal{L} = \sum_{i=1}^{n} \log \lambda(t_i \mid H_{t_i}) - M(T) - \kappa \sum_{i=1}^{n} G(T - t_i).$$
(2)

3 Latent Variable B

Using the superposition principle of Poisson processes the Hawkes process can be viewed as many different inhomogeneous Poisson processes layered up each other. After each event t_i a new Poisson process with intensity is spawned with intensity $\kappa g(t - t_i)$ between $[t_i, T]$. Under this formulation we can estimate whether an event was spawned from the background rate $(\mu(t))$ or whether it was spawned from another event during the excitation. Therefore we introduce a variable $\mathbf{B} = B_1, \ldots, B_n$ for each event t_i, \ldots, t_n that indicates who the parent is. As this can never be observed in practise it is a latent variable and instead must be estimated. The estimation probabilities can be written as

$$\Pr(B_{i} = 0 \mid t_{i}, H_{t}, \theta) = \frac{\mu(t)}{\lambda(t_{i} \mid H_{t})},$$

$$\Pr(B_{i} = j \mid t_{i}, H_{t}, \theta) = \frac{\kappa g(t_{i} - t_{j})}{\lambda(t_{i} \mid H_{t})}, \quad j = \{1, 2, \dots, i - 1\}.$$
(3)

where $B_i = 0$ indicates the event t_i came from the background rate and $B_i = j$ indicates that t_j is the parent to t_i . These latent variables allow us to split the likelihood and build a more efficient MCMC sampler.

4 Latent Variable Sampling

From the generic likelihood in (2) we introduce the branching structure from (3) to separate the likelihood

$$p(t_i, \dots t_n \mid \theta, \mathbf{B}) = e^{-\mu T} \mu^{|S_0|} \prod_{j=1}^n \left(e^{-\kappa G(T-t_j)} \kappa^{|S_j|} \prod_{t_i \in S_j} g(t_i - t_j) \right), \quad (4)$$

therefore given a sample of the latent variable \mathbf{B} we can split out the likelihood into its three components and sample new parameters from each component accordingly.

4.1 Constant Background

For all events in S_0 the likelihood is simple to write

$$p(|B) = e^{-\mu T} \mu^{|S_0|}$$

and thus the posterior distribution is

$$p(\mu \mid B) = \text{Gamma} \left(\alpha_{\mu} + \mid S_0 \mid, \beta_{\mu} + 1 \right).$$
(5)

4.2**Constant Kappa**

From the latent likelihood in (4) we can pull out the middle term that concerns κ only

$$p(\mid \theta, B, h) = \prod_{j=1}^{n} e^{-\kappa g(T-t_j)} \kappa^{|S_j|}$$
$$= \kappa^{\sum j = 1^n |S_j|} e^{-\kappa \sum_{j=1}^{n} G(T-j_j)},$$
$$\tilde{G} \equiv \sum_{j=1}^{n} G(T-j_j),$$
$$p(t_1, \dots, t_n \mid \theta, \mathbf{B}, h) = \kappa^{\sum_{j=1}^{n} |S_j|} e^{-\kappa \tilde{G}},$$
$$= \operatorname{Gamma}(\sum_{j=1}^{n} \mid S_j \mid +1, \tilde{G}),$$

we define our prior on κ as a gamma distribution also which means we can exploit the conjugacy to give the posterior distribution of κ as

$$p(\kappa \mid t_1, \dots t_n, \mathbf{B}, g) = \text{Gamma}\left(\sum_{j=1}^n \mid S_j \mid +\alpha_\kappa, \tilde{G} + \beta_\kappa\right), \quad (6)$$

where $\alpha_{\kappa}, \beta_{\kappa}$ are the prior parameters for κ . In practise $\tilde{G} \approx n$ as $G(T - t_j) = 0$ for the events that are far away from the boundary.

Exponential Kernel 4.3

We have to assume a form of g to proceed with inference. An exponential kernel is most commonly used

$$g(t \mid \beta) = \beta \exp(-\beta t),$$

we also define $\tau_i = t_i - t_j$ as the time of a child relative to its parent for all events with a parent that isn't the background. From both of these we can define the likelihood of the observations as

$$p(\mid \theta) = \prod_{j=1}^{n} \prod_{t_i \in S_j} g(t_i - t_j)$$
$$= \beta^{N_{\text{child}}} \exp\left(-\beta \sum_{i=1}^{N_{\text{child}}} \tau_i\right),$$

which again is conjugate to the gamma distribution and therefore

$$p(\beta \mid \tau_1, \dots, \tau_{N_{\text{child}}}, \mathbf{B}) = \text{Gamma}\left(N_{\text{child}} + \alpha_\beta, \sum_i^n \tau + \beta_\beta\right),$$
 (7)

where $\alpha_{\beta}, \beta_{\beta}$ are the prior parameters for the kernel.

4.4 The Algorithm

Using these posterior distributions the parameters of the Hawkes process can be updated using the following algorithm.

```
Algorithm 1: Sampling a Hawkes process using the latent variables
```

Result: *its* samples of μ , κ , β and **B while** $i \leq its$ **do** Sample new latent variables **B** using (3). ; Sample new μ using (5) ; Sample new κ using (6) ; Sample new β using (7) ;

References

- [DVJ03] D.J Daley and D Vere-Jones. An Introduction to the Theory of Point Processes. en. Probability and its Applications. New York: Springer-Verlag, 2003. ISBN: 978-0-387-95541-4. URL: http://link.springer. com/10.1007/b97277 (visited on 07/11/2016).
- [Ros16] Gordon J Ross. Bayesian Estimation of the ETAS Model for Earthquake Occurrences. 2016.